

Introduction to theory of probability and statistics

Lecture 5.

Random variable and distribution of probability

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- Concept of random variable
- Quantitative description of random variables
- Examples of probability distributions



Random variable is a function X, that attributes a real value x to a certain results of a random experiment.

$$\Omega = \{e_1, e_2, \ldots\}$$
$$X: \Omega \to R$$
$$X(e_i) = x_i \in R$$

Examples:

- 1) Coin toss: event 'head' takes a value of 1; event 'tails' 0.
- 2) Products: event 'failure' 0, well-performing 1

3) Dice: `1' - 1, `2' - 2 etc....

4) Interval [a, b] – a choice of a point of a coordinate `x' is attributed a value, e.g. sin²(3x+17) etc.



The concept of random variable

Random variable

Discrete

When the values of random variable X are isolated points on an number line

- Toss of a coin
- Transmission errors
- Faulty elements on a production line
- A number of connections coming in 5 minutes

Continuous

When the values of random variable cover all points of an interval

- Electrical current, I
- Temperature, T
- Pressure, p



- Probability distributions and probability mass functions (for discrete random variables)
- Probability density functions (for continuous variables)
- Cumulative distribution function (distribution function for discrete and continuous variables)
- Characteristic quantities (expected value, variance, quantiles, etc.)



Distribution of random variable (probability distribution for discrete variables) is a set of pairs (x_i, p_i) where x_i is a value of random variable X and p_i is a probability, that a random variable X will take a value x_i

Example 4.1

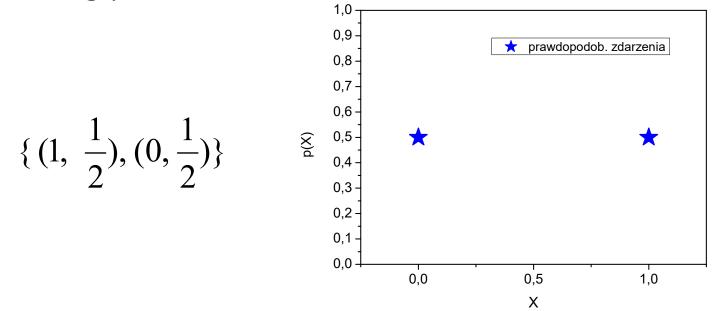
Probability mass function for a single toss of coin. Event corresponding to heads is attributed $x_1=1$; tails means $x_2=0$.

$$x_{1} = 1 \quad p(X = 1) = p(x_{1}) = \frac{1}{2}$$
$$x_{2} = 0 \quad p(X = 0) = p(x_{2}) = \frac{1}{2}$$



Example 4.1 cont.

Probability mass function for a single toss of coin is given by a set of the following pairs:



Random variable when discrete entails probability distribution also discrete.



Probability density function

Probability function is introduced for continuous variables; it is related to probability in the following way:

$$f(x)dx \equiv P(x \le X < x + dx)$$

Properties of probability density function:

1.
$$f(x) \ge 0$$

2. $f(x)$ is normalized $\int_{-\infty}^{+\infty} f(x) dx = 1$

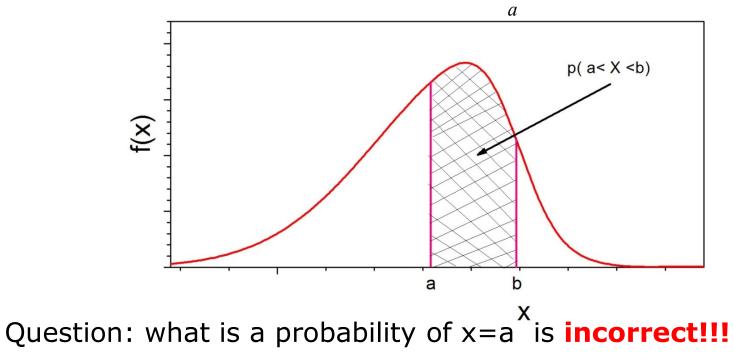
3. f(x) has a measure of 1/x



Probability density function

Directly from a definition of probability density function f(x) we get a formula of calculating the probability that the random variable will assume a value within an interval of [a,b]:

$$P(a < X < b) = \int f(x) dx$$

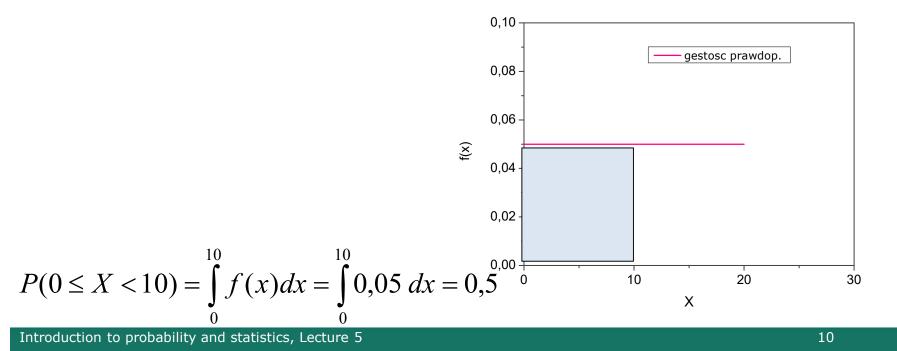




Probability density function

Example 4.2

Let the continuous random variable X denote the current measured in a thin copper wire in mA. Assume that the range of X is [0, 20 mA], and assume that the probability density function of X is f(x)=0,05 for $0 \le x \le 20$. What is the probability that a current measured is less than 10 mA.

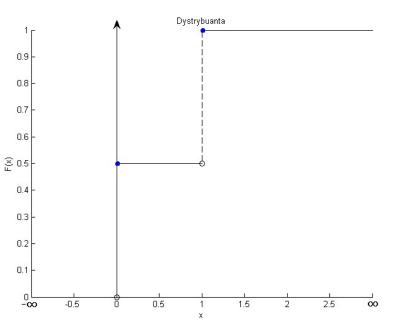


Quantitative description of random variables

• Cumulative distribution function (CDF) F(x) is a probability of an event that the random variable X will assume a value smaller than or equal to x (at most x) $F(x) = P(X \le x)$

CDF of coin toss:

$$F(x=0) = P(X \le 0) = \frac{1}{2}$$
$$F(x=1) = P(X \le 1) = 1$$





Properties of CDF

1. $0 \le F(x) \le 1$ 2. $F(-\infty) = 0$ 3. $F(+\infty) = 1$ 4. $x \le y \implies F(x) \le F(y)$

non-decreasing function

5. F(x) has no unit

6.
$$f(x) = \frac{dF(x)}{dx}$$

Relationship between cumulative distribution function and probability density (for continuous variable)



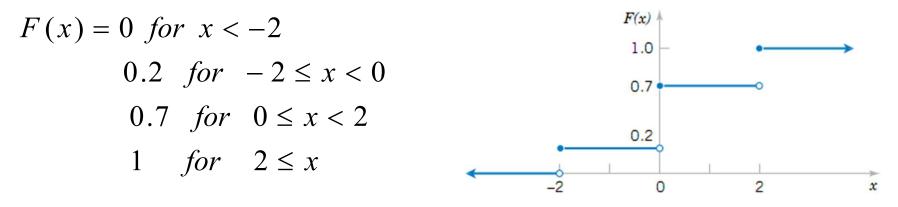
CDF of discrete variable

$$F(x) = P(X \le x) = \sum_{x_i \le x} f(x_i)$$

f (x_i) – probability mass function

Example 4.3

Determine probability mass function of X from the following cumulative distribution function F(x)



From the plot, the only points to receive $f(x) \neq 0$ are -2, 0, 2.

$$f(-2) = 0.2 - 0 = 0.2$$
 $f(0) = 0.7 - 0.2 = 0.5$ $f(2) = 1.0 - 0.7 = 0.3$

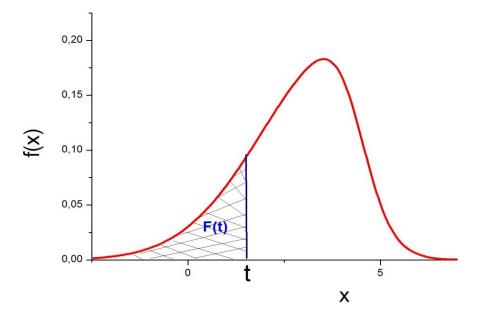


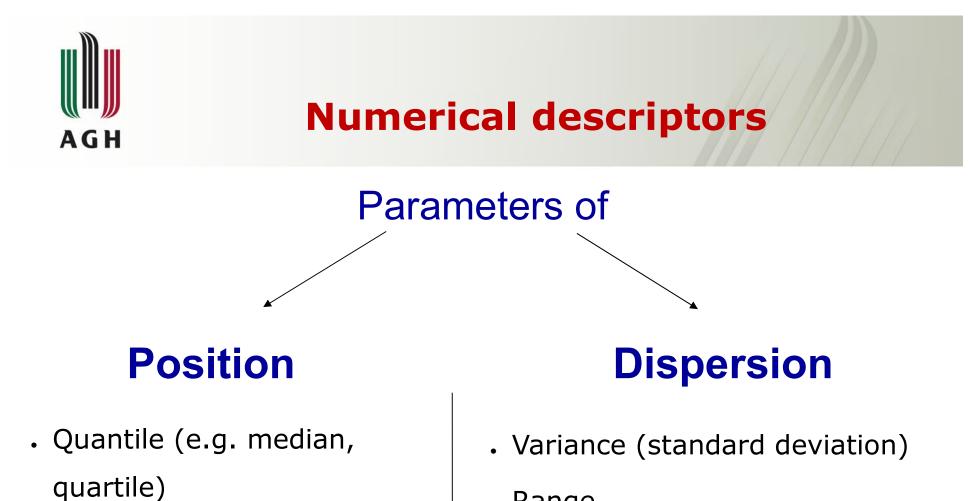
CDF for continuous variable

4

$$F(t) = P(X \le t) = \int_{-\infty}^{t} f(x) \, dx$$

Cumulative distribution function F(t) of continuous variable is a nondecreasing continuous function and can be calculated as an area under density probability function f(x) over an interval from - ∞ to t.





- . Mode
- . Expected value (average)

Range



Numerical descriptors

Quantile x_q represents a value of random variable for which the cumulative distribution function takes a value of q.

$$F(x_q) = P(X \le x_q) = \int_{-\infty}^{x_q} f(u) \, du = q$$

Median i.e. $x_{0.5}$ is the most frequently used quantile.

In example 4.2 current I=10 mA is a median of distribution.

Example 4.4

For a discrete distribution : 19, 21, 21, 21, 22, 22, 23, 25, 26, 27 median is 22 (middle value or arithmetic average of two middle values)



Numerical descriptors

Mode represents the most frequently occurring value of random variable (x at which probability distribution attains a maximum)

Unimodal distribution has one mode (multimodal distributions – more than one mode)

In example 4.4: x_k = 19, 21, 21, 21, 22, 22, 23, 25, 26, 27 mode equals to 21 (which appears 3 times, i.e., the most frequently)

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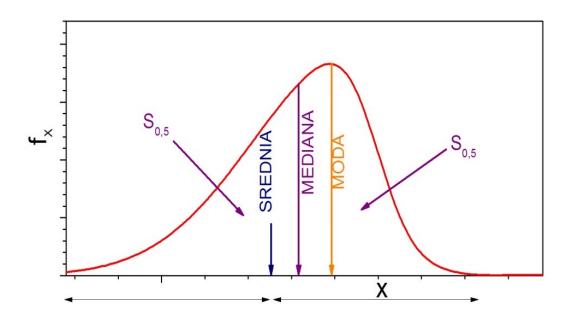
Average value

Arithmetic average:

 \boldsymbol{x}_i - belongs to a set of n – elements

$$\overline{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$$

In example 4.4: $x_i = 19, 21, 21, 21, 22, 22, 23, 25, 26, 27$, the arithmetic average is 22.7





Arithmetic average

n

Many elements having the same value, we divide the set into classes containing n_k identical elements

Example 4.5

X _k	n _k	f _k
10.2	1	0.0357
12.3	4	0.1429
12.4	2	0.0714
13.4	8	0.2857
16.4	4	0.1429
17.5	3	0.1071
19.3	1	0.0357
21.4	2	0.0714
22.4	2	0.0714
25.2	1	0.0357
Sum	28	

$\overline{x} = \frac{k}{k}$	$\sum_{k=1}^{p} n_k x_k$	$-=\sum_{k=1}^{p}f_{k}x_{k}$	
	n	$\sum_{k=1}^{J} \int k^{k} v_{k}$	ł

where: $f_k = \frac{n_k}{n}$, p - number of classes $(p \le n)$ Normalization condition $\sum_k f_k = 1$ $\overline{x} = x_1 \cdot f_1 + x_2 \cdot f_2 + \ldots + x_n \cdot f_n =$ $= 10.2 \cdot 0.04 + 12.3 \cdot 0.14 + \ldots + 25.2 \cdot 0.04$ $\overline{x} = 15.77$



Moment of the order k with respect to x_0

$$m_k(x_0) \equiv \sum_i (x_i - x_0)^k p(x_i)$$
 for discrete variables

$$m_k(x_0) \equiv \int (x - x_0)^k f(x) dx$$
 for continuous variables

The most important are the moments calculated with respect to $x_0=0$ (m_k) and $X_0=m_1$ the first moment (m_1 is called the expected value) – these are central moments μ_k .



Expected value

Symbols: m_1 , E(X), μ , \overline{x} , \hat{x}

$$E(X) = \sum_{i} x_{i} p_{i}$$
 for discrete variables

$$E(X) \equiv \int x f(x) dx \qquad \text{for continuous variables}$$



Properties of E(X)

E(X) is a linear operator, i.e.:

1.
$$E(\sum_{i} C_{i}X_{i}) = \sum_{i} C_{i}E(X_{i})$$

In a consequence:

$$E(C) = C$$

$$E(CX) = CE(X)$$

$$E(X_1 + X_2) = E(X_1) + E(X_2)$$

2. For independent variables $X_{1_i} X_{2_i} \dots X_n$ $E(\prod_i X_i) = \prod_i E(X_i)$

Variables are independent when:

$$f(X_1, X_2, ..., X_n) = f_1(X_1) f_2(X_2) \cdot ... \cdot f_n(X_n)$$



Properties of E(X)

3. For a function of X; Y = Y(X) the expected value E(Y) can be found on the basis of distribution of variable X without necessity of looking for distribution of f(y)

$$E(Y) = \sum_{i} y(x_{i})p_{i}$$
 for discrete variables
$$E(Y) \equiv \int y(x) f(x) dx$$
 for continuous variables

Any moment $m_k(x_0)$ can be treated as an expected value of a function $Y(X)=(X-x_0)^k$

$$m_k(x_0) \equiv \int (x - x_0)^k f(x) \, dx = E((x - x_0)^k)$$



Variance

VARIANCE (dispersion) symbols: $\sigma^2(X)$, var(X), V(X), D(X). *Standard deviation* $\sigma(x)$

$$\sigma^{2}(X) \equiv \sum_{i} p_{i} (x_{i} - E(X))^{2}$$
 for discrete variables

$$\sigma^{2}(X) \equiv \int f(x)(x - E(X)^{2} dx \quad \text{for continuous variables}$$

Variance (or the standard deviation) is a measure of scatter of random variables around the expected value.

$$\sigma^2(X) = E(X^2) - E^2(X)$$



Properties of $\sigma^2(X)$

Variance can be calculated using expected values only:

1.
$$\sigma^2(X) = E(X^2) - E^2(X)$$

In a consequence we get:

$$\sigma^{2}(C) = 0$$

$$\sigma^{2}(CX) = C^{2} \sigma^{2}(X)$$

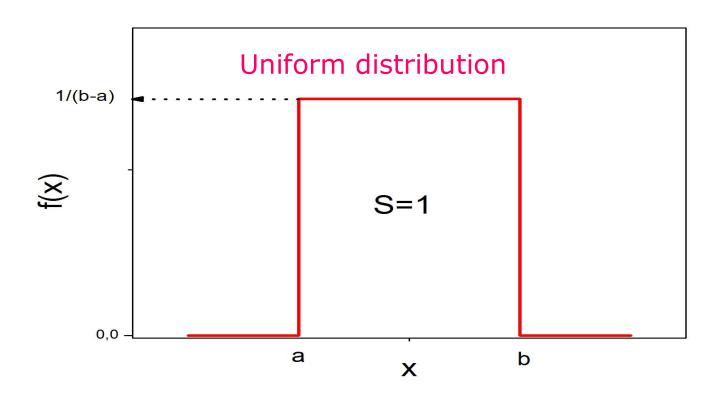
$$\sigma^{2}(C_{1}X + C_{2}) = C_{1}^{2} \sigma^{2}(X)$$

2. For independent variables $X_{1,} X_{2,} \dots X_{n}$

$$\sigma^2(\sum_i C_i X_i) = \sum_i C_i^2 \sigma^2(X)$$

UNIFORM DISTRIBUTION

$$f(x) = \frac{1}{b-a} \quad \mu = EX = \frac{a+b}{2} \quad \sigma^2 = \frac{(b-a)^2}{12}$$
$$a \le x \le b$$



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Interpretation of variance results from Czebyszew theorem:

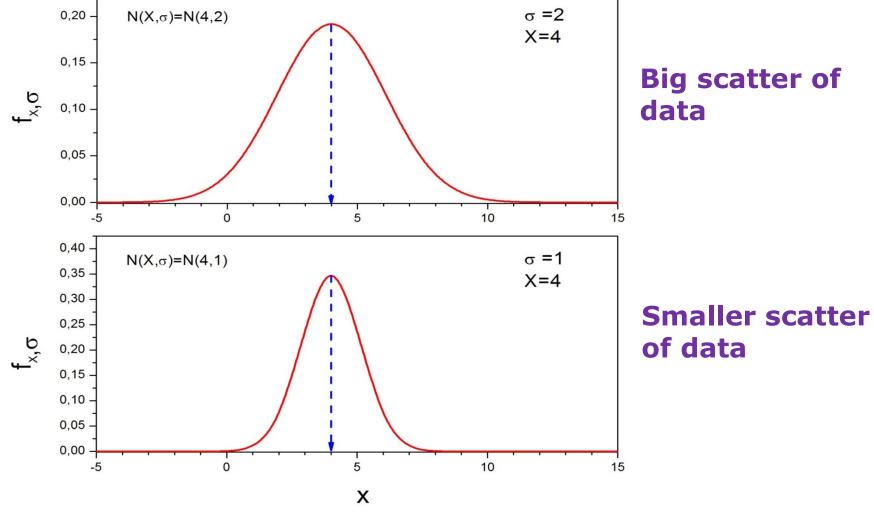
$$P(|X - E(X)| \ge a.\sigma(X)) \le \frac{1}{a^2}$$

Theorem:

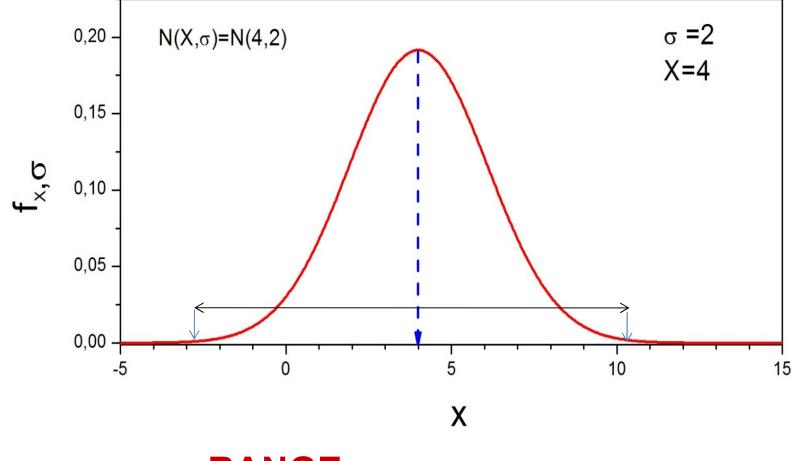
Probability of the random variable X to be shifted from the expected value E(X) by a-times standard deviation is smaller or equal to $1/a^2$

This theorem is valid for all distributions that have a variance and the expected value. Number a is any positive real value.









RANGE = x_{max} - x_{min}



Practical ways of calculating variance

Variance of n-element sample:

$$s^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (x_{i} - \overline{x})^{2} \qquad s^{2} = \frac{1}{n-1} \left[\sum_{i=1}^{n} x_{i}^{2} - \frac{(\sum_{i=1}^{n} x_{i})^{2}}{n} \right]$$
$$\overline{x} - average$$

Variance of N-element population :

$$\sigma^{2} = \frac{1}{N} \sum_{i=1}^{N} (x_{i} - \mu)^{2}$$
$$\mu - expected \quad value$$



Practical ways of calculating standard deviation

Standard deviation of sample (or: standard uncertainty):

$$s = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} (x_i - \overline{x})^2}$$

Standard deviation (population):

$$\sigma = \sqrt{\frac{1}{N} \sum_{i=1}^{N} (x_i - \mu)^2}$$



Two-point distribution (zero-one), e.g. coin toss, head = failure x=0, tail = success x=1, p - probability of success, its distribution:

Binomial (Bernoulli)

$$p_{k} = {\binom{n}{k}} \cdot p^{k} (1-p)^{n-k} , k = 0, 1, ..., n$$

where $0 ; <math>X = \{0, 1, 2, ..., k\}$ k – number of successes when n-times sampled with replacement

For k=1 two-point distribution



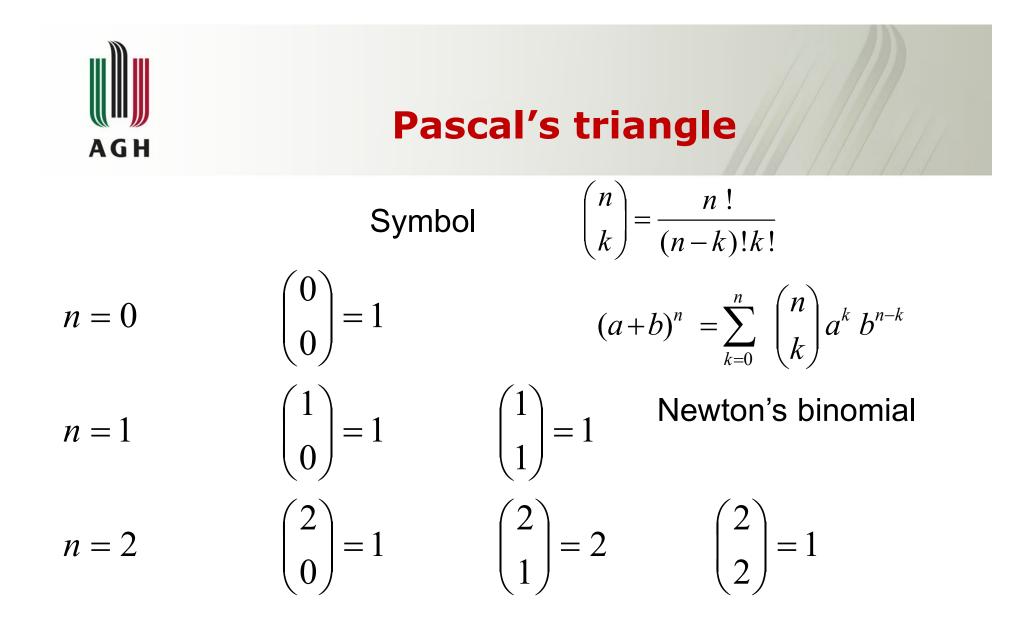
Binomial distribution assumptions

Random experiment consists of *n* Bernoulli trials :

- 1. Each trial is independent of others.
- 2. Each trial can have only two results: "success" and "failure" (binary!).
- 3. Probability of success *p* is <u>constant.</u>

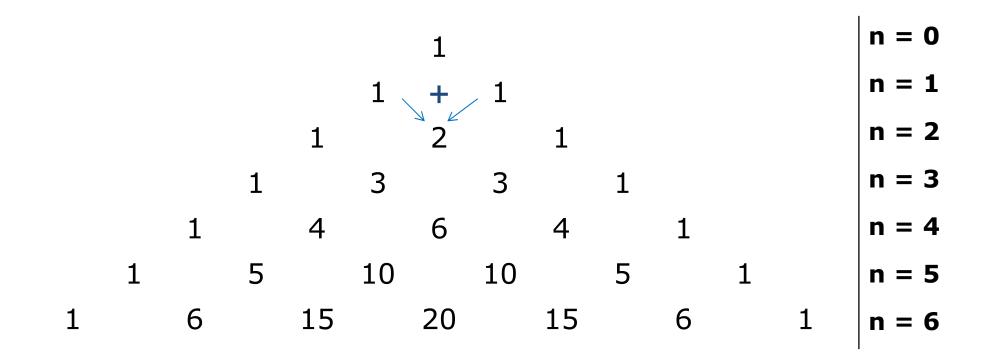
Probability p_k of an event that random variable X will be equal to the number of k-successes at n trials.

$$p_{k} = {\binom{n}{k}} \cdot p^{k} (1-p)^{n-k} , k = 0, 1, ..., n$$





Pascal's triangle





Bernoulli distribution

Example 4.6

Probability that in a company the daily use of water will not exceed a certain level is p=3/4. We monitor a use of water for 6 days.

Calculate a probability the daily use of water will not exceed the set-up limit in 0, 1, 2, ..., 6 consecutive days, respectively.

Data:

$$p = \frac{3}{4}$$
 $q = \frac{1}{4}$ $N = 6$ $k = 0, 1, \dots, 6$



$$k = 0 \qquad P(k = 0) = \begin{pmatrix} 6\\0 \end{pmatrix} \cdot \left(\frac{3}{4}\right)^0 \cdot \left(\frac{1}{4}\right)^6$$

$$k = 1 \qquad P(k = 1) = \begin{pmatrix} 6\\0 \end{pmatrix} \cdot \left(\frac{3}{4}\right)^1 \cdot \left(\frac{1}{4}\right)^5$$

$$k = 1 P(k = 1) = \begin{pmatrix} 0 \\ 1 \end{pmatrix} \cdot \begin{pmatrix} \frac{3}{4} \end{pmatrix} \cdot \begin{pmatrix} \frac{1}{4} \end{pmatrix}$$

$$k = 2 \qquad P(k = 2) = {\binom{6}{2}} \cdot \left(\frac{3}{4}\right)^2 \cdot \left(\frac{1}{4}\right)^4$$

$$k = 3$$
 $P(k = 3) = \begin{pmatrix} 6 \\ 3 \end{pmatrix} \cdot \left(\frac{3}{4}\right)^3 \cdot \left(\frac{1}{4}\right)^3$

$$k = 4 \qquad P(k = 4) = {\binom{6}{4}} \cdot {\left(\frac{3}{4}\right)}^4 \cdot {\left(\frac{1}{4}\right)}^2$$

$$k = 5 \qquad P(k = 5) = {\binom{6}{5}} \cdot {\left(\frac{3}{4}\right)}^5 \cdot {\left(\frac{1}{4}\right)}^1$$

$$k = 6 \qquad P(k = 6) = {\binom{6}{6}} \cdot \left(\frac{3}{4}\right)^6 \cdot \left(\frac{1}{4}\right)^0$$



$$k = 0 \qquad P(0) = 1 \cdot 1 \cdot \frac{1}{4^6} \cong 0.00024$$

$$k = 1 \qquad P(1) = 6 \cdot \frac{3}{4} \cdot \frac{1}{4^5} = \frac{6 \cdot 3}{4^6} = 18 \cdot P(0) \cong 0.004$$

$$k = 2 \qquad P(2) = 15 \cdot \left(\frac{3}{4}\right)^2 \cdot \frac{1}{4^4} = \frac{15 \cdot 9}{4^6} = 135 \cdot P(0) \cong 0.033$$

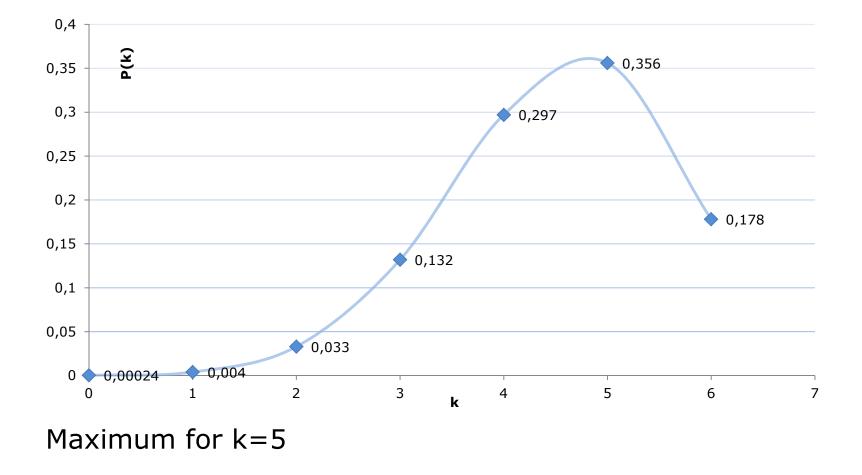
$$k = 3 \qquad P(3) = 20 \cdot \left(\frac{3}{4}\right)^3 \cdot \frac{1}{4^3} = \frac{20 \cdot 9 \cdot 3}{4^6} = 540 \cdot P(0) \cong 0.132$$

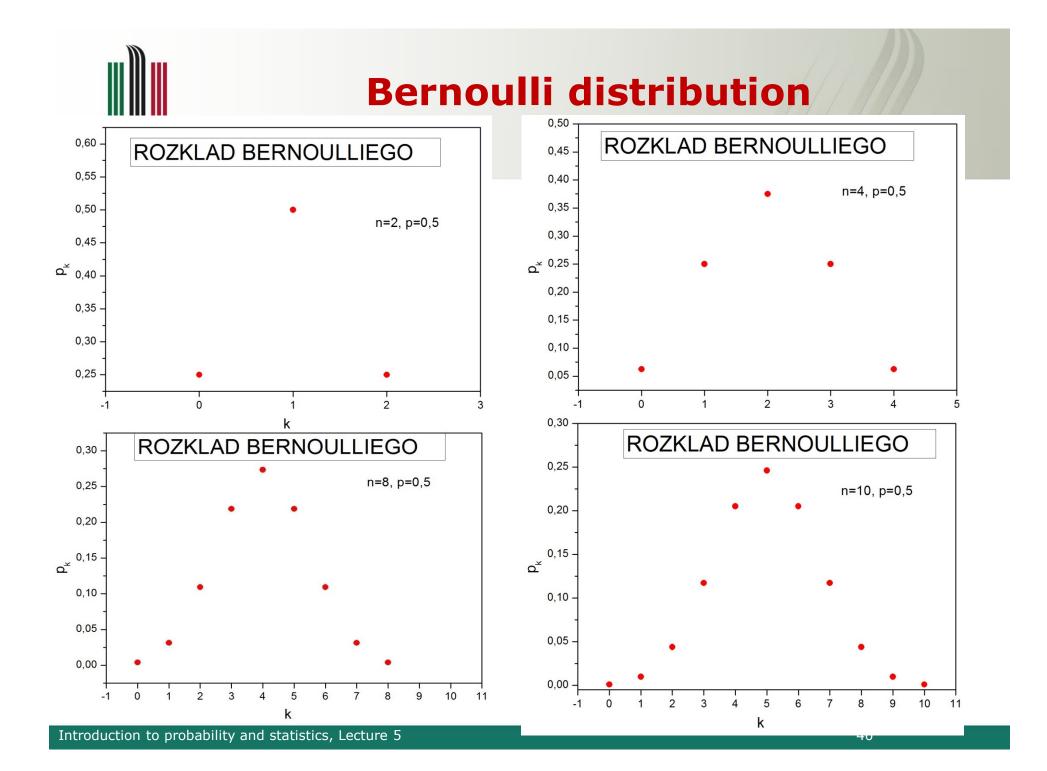
$$k = 4 \qquad P(4) = 15 \cdot \left(\frac{3}{4}\right)^4 \cdot \frac{1}{4^2} = \frac{15 \cdot 9 \cdot 9}{4^6} = 1215 \cdot P(0) \cong 0.297$$

$$k = 5 \qquad P(5) = 6 \cdot \left(\frac{3}{4}\right)^5 \cdot \frac{1}{4^1} = \frac{6 \cdot 9 \cdot 9 \cdot 3}{4^6} = 1458 \cdot P(0) \cong 0.356$$

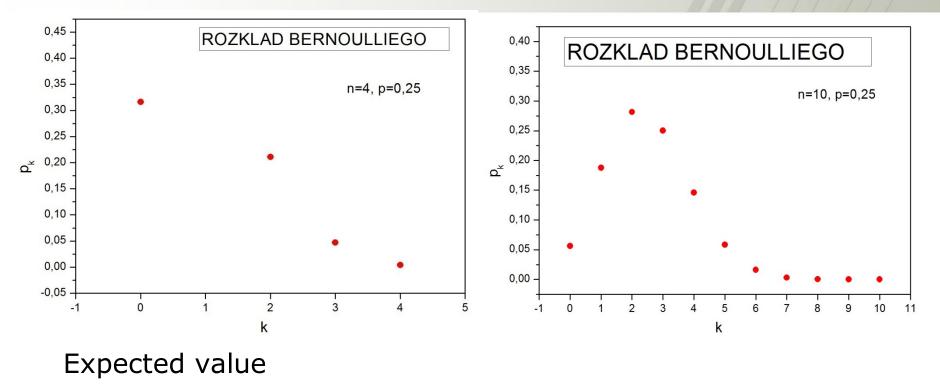
$$k = 6 \qquad P(6) = 1 \cdot \left(\frac{3}{4}\right)^6 \cdot \frac{1}{4^0} = \frac{9 \cdot 9 \cdot 9}{4^6} = 729 \cdot P(0) \cong 0.178$$











 $E(X) = \mu = np$

Variance

$$V(X) = \sigma^2 = np(1-p)$$



Errors in transmission

Example 4.7

Digital channel of information transfer is prone to errors in single bits. Assume that the probability of single bit error is p=0.1

Consecutive errors in transmissions are independent. Let X denote the random variable, of values equal to the number of bits in error, in a sequence of 4 bits.

E - bit error, O - no error
OEOE corresponds to X=2; for EEOO - X=2 (order does not matter)



Errors in transmission

Example 4.7 cd

For X=2 we get the following results: {EEOO, EOEO, EOOE, OEEO, OEOE, OOEE}

What is a probability of P(X=2), i.e., two bits will be sent with error?

Events are independent, thus $P(EEOO)=P(E)P(O)P(O)=(0.1)^2 (0.9)^2 = 0.0081$

Events are mutually exhaustive and have the same probability, hence P(X=2)=6 P(EEOO)= 6 $(0.1)^2 (0.9)^2 = 6 (0.0081)=0.0486$



Errors in transmission

Example 4.7 continued

$$\binom{4}{2} = \frac{4!}{(2)!2!} = 6$$

Therefore, $P(X=2)=6 (0.1)^2 (0.9)^2$ is given by Bernoulli distribution

$$P(X = x) = \binom{4}{x} \cdot p^{x} (1 - p)^{4 - x}, x = 0, 1, 2, 3, 4, p = 0.1$$

$$P(X = 0) = 0,6561$$

$$P(X = 1) = 0,2916$$

$$P(X = 2) = 0,0486$$

$$P(X = 3) = 0,0036$$

$$P(X = 4) = 0,0001$$

0

1

2

3

4



Poisson's distribution

We introduce a parameter $\lambda = pn (E(X) = \lambda)$

$$P(X = x) = \binom{n}{x} \cdot p^{x} (1 - p)^{n - x} = \binom{n}{x} \left(\frac{\lambda}{n}\right)^{x} \left(1 - \frac{\lambda}{n}\right)^{n - x}$$

Let us assume that n increases while p decreases, but λ =pn remains constant. Bernoulli distribution changes to Poisson's distribution.

$$\lim_{n \to \infty} P(X = x) = \lim_{n \to \infty} {n \choose x} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x} = \frac{e^{-\lambda}\lambda^x}{x!}$$



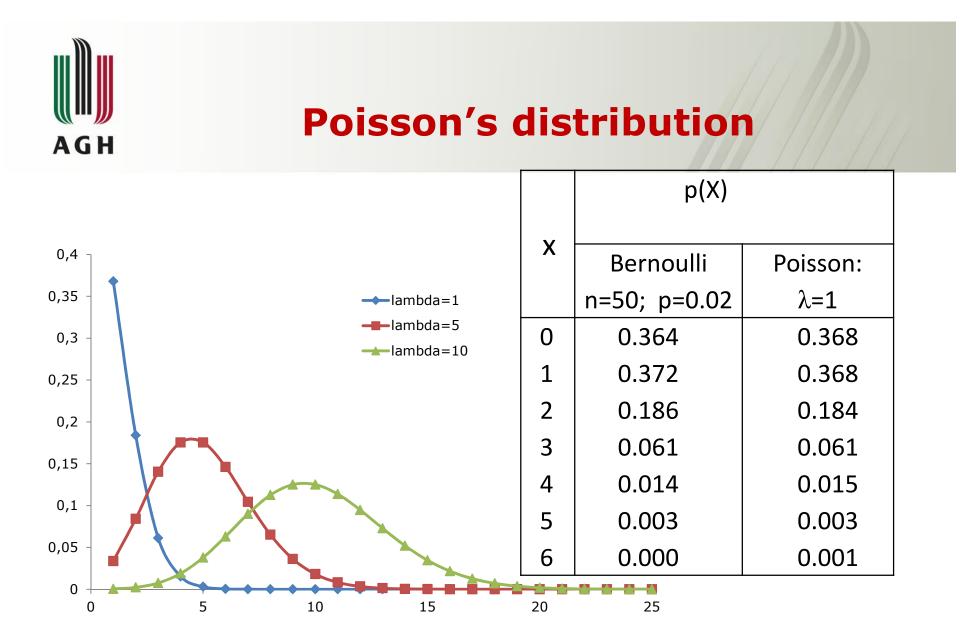
Poisson's distribution

It is one of the rare cases where expected value equals to variance:

$$E(X) = np = \lambda$$

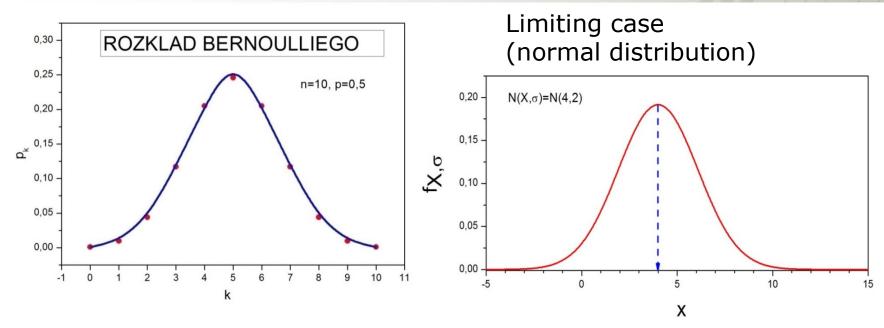
Why?

$$V(X) = \sigma^{2} = \lim_{n \to \infty, p \to 0} (np - np^{2}) = np = \lambda$$



(x- integer, infinite; $x \ge 0$) For big n Bernoulli distribution resembles Poisson's distribution





The most widely used model for the distribution of random variable is a **normal distribution**.

Central limit theorem formulated in 1733 by De Moivre

Whenever a random experiment is replicated, the random variable that equals the average (or total) result over the replicas tends to have a normal distribution as the number of replicas becomes large.



A random variable X with probability density function f(x):

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(x-\mu^2)}{2\sigma^2}\right], \text{ where } -\infty < x < +\infty$$

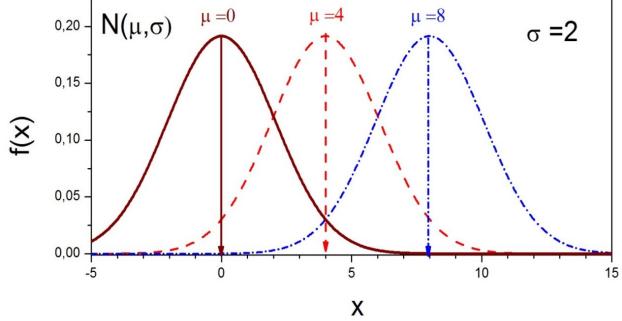
is a **normal random variable** with two parameters:

$$-\infty < \mu < +\infty, \sigma > 1$$

We can show that $E(X) = \mu$ and $V(X) = \sigma^2$

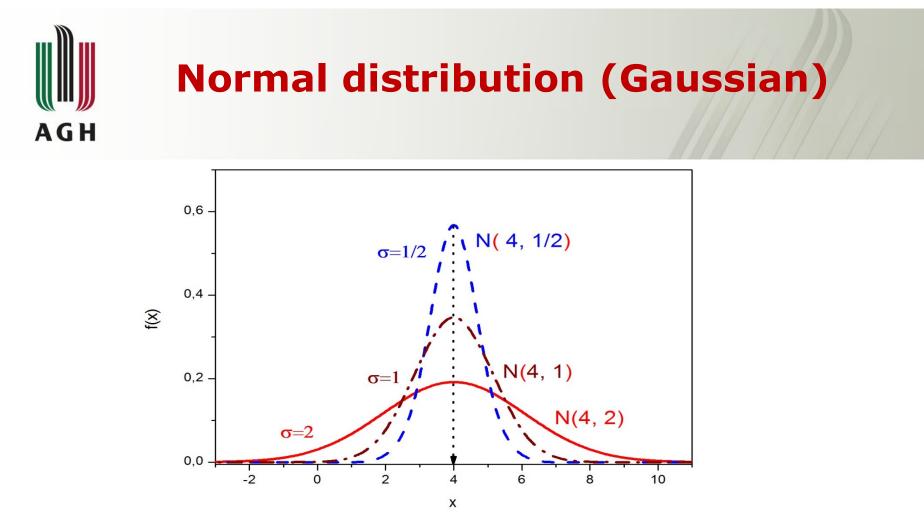
Notation $N(\mu,\sigma)$ is used to denote this distribution





Expected value, maximum of density probability (mode) and median overlap ($x=\mu$). Symmetric curve (Gaussian curve is bell shaped).

Variance is a measure of the width of distribution. At $x=+\sigma$ and $x=-\sigma$ there are the inflection points of N(0, σ).



Is used in experimental physics and describes distribution of random errors. Standard deviation σ is a measure of random uncertainty. Measurements with larger σ correspond to bigger scatter of data around the average value and thus have **less precision**.

Standard normal distribution

A normal random variable Z with probability density N(z):

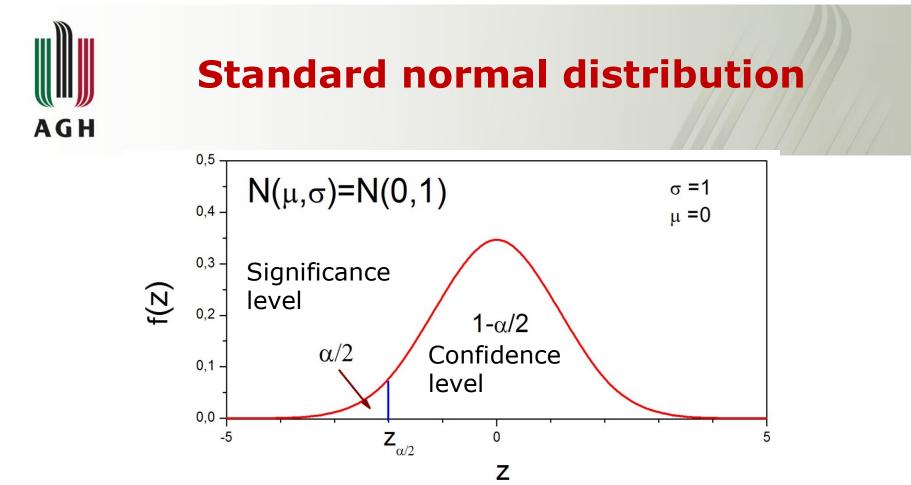
$$N(z) = \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{z^2}{2}\right] , where - \infty < z < +\infty$$

is called a **standard normal random variable**

N(0,1)
$$E(Z) = 0, V(Z) = 1$$

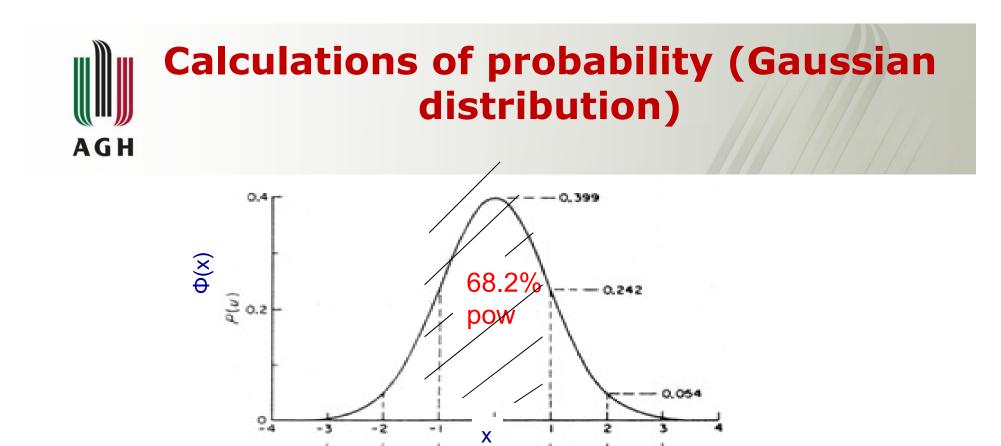
Definition of standard normal variable

$$Z = \frac{X - \mu}{\sigma}$$



Advantages of standardization:

- Tables of values of probability density and CDF can be constructed for N(0,1). A new variable of the N(μ , σ) distribution can be created by a simple transformation X= σ *Z+ μ
- By standardization we shift all original random variables to the region close to zero and we rescale the x-axis. The unit changes to standard deviation. Therefore, we can compare different distribution.



 $(-\sigma, +\sigma)$

 $(-2\sigma, + 2\sigma)$

 $(-3\sigma, + 3\sigma)$

 $P(\mu-\sigma < X < \mu+\sigma) = 0,6827$ (about 2/3 of results) $P(\mu-2\sigma < X < \mu+2\sigma) = 0,9545$ $P(\mu-2\sigma < X < \mu+2\sigma) = 0,9973$ (almost all)