## Introduction to theory of probability and statistics

## Lecture 5.

## Random variable and distribution of probability

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## Outline:

. Concept of random variable

- Quantitative description of random variables
. Examples of probability distributions


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## The concept of random variable

Random variable is a function $X$, that attributes a real value $x$ to a certain results of a random experiment.

$$
\begin{aligned}
& \Omega=\left\{e_{1}, e_{2}, \ldots\right\} \\
& X: \Omega \rightarrow R \\
& X\left(e_{i}\right)=x_{i} \in R
\end{aligned}
$$

Examples:

1) Coin toss: event 'head' takes a value of 1 ; event 'tails' - 0.
2) Products: event 'failure' - 0, well-performing - 1
3) Dice: '1' - 1, '2' - 2 etc....
4) Interval [a, b]- a choice of a point of a coordinate ' $x$ ' is attributed a value, e.g. $\sin ^{2}(3 x+17)$ etc. ....

## The concept of random variable

## Random variable



## Discrete

When the values of random variable $X$ are isolated points on an number line

- Toss of a coin
- Transmission errors
- Faulty elements on a production line
- A number of connections coming in 5 minutes

Continuous
When the values of random variable cover all points of an interval

- Electrical current, I
- Temperature, $T$
- Pressure, p


## Quantitative description of random variables

- Probability distributions and probability mass functions (for discrete random variables)
- Probability density functions (for continuous variables)
- Cumulative distribution function (distribution function for discrete and continuous variables)
- Characteristic quantities (expected value, variance, quantiles, etc.)


## Distribution of random variable

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Distribution of random variable (probability distribution for discrete variables) is a set of pairs ( $x_{i}, p_{i}$ ) where $x_{i}$ is a value of random variable $X$ and $p_{i}$ is a probability, that a random variable $X$ will take a value $x_{i}$

## Example 4.1

Probability mass function for a single toss of coin.
Event corresponding to heads is attributed $x_{1}=1$; tails means $x_{2}=0$.

$$
\begin{aligned}
& x_{1}=1 \quad p(X=1)=p\left(x_{1}\right)=\frac{1}{2} \\
& x_{2}=0 \quad p(X=0)=p\left(x_{2}\right)=\frac{1}{2}
\end{aligned}
$$

## Distribution of random variable

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## Example 4.1 cont.

Probability mass function for a single toss of coin is given by a set of the following pairs:

Random variable when discrete entails probability distribution also discrete.

## Probability density function

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Probability function is introduced for continuous variables; it is related to probability in the following way:

$$
f(x) d x \equiv P(x \leq X<x+d x)
$$

Properties of probability density function:

$$
\text { 1. } f(x) \geq 0
$$

2. $f(x)$ is normalized $\int_{-\infty}^{+\infty} f(x) d x=1$
3. $f(x)$ has a measure of $1 / x$

## Probability density function

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Directly from a definition of probability density function $f(x)$ we get a formula of calculating the probability that the random variable will assume a value within an interval of $[\mathrm{a}, \mathrm{b}]$ :

$$
P(a<X<b)=\int f(x) d x
$$



Question: what is a probability of $\mathrm{x}=\mathrm{a}^{\mathrm{x}}$ is incorrect!!!

## Probability density function

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## Example 4.2

Let the continuous random variable $X$ denote the current measured in a thin copper wire in mA. Assume that the range of $X$ is $[0,20 \mathrm{~mA}]$, and assume that the probability density function of $X$ is $f(x)=0,05$ for $0 \leq x \leq 20$. What is the probability that a current measured is less than 10 mA .


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## Quantitative description of random variables

- Cumulative distribution function (CDF) $F(x)$ is a probability of an event that the random variable $X$ will assume a value smaller than or equal to $x$ (at most $x$ )

$$
F(x)=P(X \leq x)
$$

## Example 4.1 cont.

CDF of coin toss:

$$
\begin{gathered}
F(x=0)=P(X \leq 0)=\frac{1}{2} \\
F(x=1)=P(X \leq 1)=1
\end{gathered}
$$



## Properties of CDF

1. $0 \leq F(x) \leq 1$
2. $F(-\infty)=0$
3. $F(+\infty)=1$
4. $x \leq y \Rightarrow F(x) \leq F(y)$
non-decreasing function
5. $F(x)$ has no unit
6. $f(x)=\frac{d F(x)}{d x} \begin{aligned} & \begin{array}{l}\text { Relationship between cumulative } \\ \text { distribution function and probability } \\ \text { density (for continuous variable) }\end{array}\end{aligned}$

## CDF of discrete variable

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$$
F(x)=P(X \leq x)=\sum_{x_{i} \leq x} f\left(x_{i}\right)
$$

$f\left(x_{i}\right)$ - probability mass function

## Example 4.3

Determine probability mass function of $X$ from the following cumulative distribution function $F(x)$

$$
\begin{aligned}
& F(x)= 0 \text { for } x<-2 \\
& 0.2 \text { for }-2 \leq x<0 \\
& 0.7 \text { for } 0 \leq x<2 \\
& 1 \text { for } 2 \leq x
\end{aligned}
$$



From the plot, the only points to receive $f(x) \neq 0$ are $-2,0,2$.
$f(-2)=0.2-0=0.2 \quad f(0)=0.7-0.2=0.5 \quad f(2)=1.0-0.7=0.3$

## CDF for continuous variable

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$$
F(t)=P(X \leq t)=\int_{-\infty}^{t} f(x) d x
$$

Cumulative distribution function $F(t)$ of continuous variable is a nondecreasing continuous function and can be calculated as an area under density probability function $f(x)$ over an interval from - $\infty$ to $t$.


## Numerical descriptors

## Parameters of

## Position

- Quantile (e.g. median, quartile)
. Mode
- Expected value (average)


## Dispersion

- Variance (standard deviation)
- Range


## Numerical descriptors

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Quantile $x_{q}$ represents a value of random variable for which the cumulative distribution function takes a value of $q$.

$$
F\left(x_{q}\right)=P\left(X \leq x_{q}\right)=\int_{-\infty}^{x_{q}} f(u) d u=q
$$

Median i.e. $\mathrm{x}_{0.5}$ is the most frequently used quantile.
In example 4.2 current $\mathrm{I}=10 \mathrm{~mA}$ is a median of distribution.

## Example 4.4

For a discrete distribution : 19, 21, 21, 21, 22, 22, 23, 25, 26, 27 median is 22 (middle value or arithmetic average of two middle values)

## Numerical descriptors

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Mode represents the most frequently occurring value of random variable ( $x$ at which probability distribution attains a maximum)

Unimodal distribution has one mode (multimodal distributions more than one mode)

In example 4.4: $x_{k}=19,21,21,21,22,22,23,25,26,27$ mode equals to 21 (which appears 3 times, i.e., the most frequently)

## Average value

## Arithmetic average:

$\mathrm{x}_{\mathrm{i}}$ - belongs to a set of n - elements

$$
\bar{x}=\frac{1}{n} \sum_{i=1}^{n} x_{i}
$$

In example 4.4: $x_{i}=19,21,21,21,22,22,23,25,26,27$, the arithmetic average is 22.7


## Arithmetic average

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Many elements having the same value, we divide the set into classes containing $n_{k}$ identical elements

## Example 4.5

| $\mathrm{x}_{\mathrm{k}}$ | $\mathrm{n}_{\mathrm{k}}$ | $\mathrm{f}_{\mathrm{k}}$ |
| ---: | ---: | :---: |
| 10.2 | 1 | 0.0357 |
| 12.3 | 4 | 0.1429 |
| 12.4 | 2 | 0.0714 |
| 13.4 | 8 | 0.2857 |
| 16.4 | 4 | 0.1429 |
| 17.5 | 3 | 0.1071 |
| 19.3 | 1 | 0.0357 |
| 21.4 | 2 | 0.0714 |
| 22.4 | 2 | 0.0714 |
| 25.2 | 1 | 0.0357 |
| Sum | 28 |  |

$$
\bar{x}=\frac{\sum_{k=1}^{p} n_{k} x_{k}}{n}=\sum_{k=1}^{p} f_{k} x_{k}
$$

$$
\text { where: } f_{k}=\frac{n_{k}}{n}, p-n u m b e r \text { of classes }(p \leq n)
$$

Normalization condition $\quad \sum_{k} f_{k}=1$

$$
\begin{gathered}
\bar{x}=x_{1} \cdot f_{1}+x_{2} \cdot f_{2}+\ldots+x_{n} \cdot f_{n}= \\
=10.2 \cdot 0.04+12.3 \cdot 0.14+\ldots+25.2 \cdot 0.04 \\
\bar{x}=15.77
\end{gathered}
$$

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## Moments of distribution functions

Moment of the order $k$ with respect to $x_{0}$

$$
\begin{aligned}
& m_{k}\left(x_{0}\right) \equiv \sum_{i}\left(x_{i}-x_{0}\right)^{k} p\left(x_{i}\right) \quad \text { for discrete variables } \\
& m_{k}\left(x_{0}\right) \equiv \int\left(x-x_{0}\right)^{k} f(x) d x \quad \text { for continuous variables }
\end{aligned}
$$

The most important are the moments calculated with respect to $x_{0}=0\left(m_{k}\right)$ and $X_{0}=m_{1}$ the first moment ( $m_{1}$ is called the expected value) - these are central moments $\mu_{\mathrm{k}}$.

## Expected value

Symbols: $\quad m_{1}, E(X), \mu, \bar{x}, \hat{x}$

$$
E(X)=\sum_{i} x_{i} p_{i} \quad \text { for discrete variables }
$$

$$
E(X) \equiv \int x f(x) d x \quad \text { for continuous variables }
$$

## Properties of $E(X)$

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$E(X)$ is a linear operator, i.e.:
1.

$$
E\left(\sum_{i} C_{i} X_{i}\right)=\sum_{i} C_{i} E\left(X_{i}\right)
$$

In a consequence:

$$
\begin{aligned}
& \mathrm{E}(\mathrm{C})=\mathrm{C} \\
& \mathrm{E}(\mathrm{CX})=\mathrm{CE}(\mathrm{X}) \\
& \mathrm{E}\left(\mathrm{X}_{1}+\mathrm{X}_{2}\right)=\mathrm{E}\left(\mathrm{X}_{1}\right)+\mathrm{E}\left(\mathrm{X}_{2}\right)
\end{aligned}
$$

2. For independent variables $X_{1}, X_{2}, \ldots X_{n}$

$$
E\left(\prod_{i} X_{i}\right)=\prod_{i} E\left(X_{i}\right)
$$

Variables are independent when:

$$
f\left(X_{1}, X_{2}, \ldots, X_{n}\right)=f_{1}\left(X_{1}\right) f_{2}\left(X_{2}\right) \cdot \ldots \cdot f_{n}\left(X_{n}\right)
$$

## Properties of $\mathbf{E}(\mathbf{X})$

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3. For a function of $X ; Y=Y(X)$ the expected value $E(Y)$ can be found on the basis of distribution of variable $X$ without necessity of looking for distribution of $f(y)$

$$
\begin{array}{ll}
E(Y)=\sum_{i} y\left(x_{i}\right) p_{i} & \text { for discrete variables } \\
E(Y) \equiv \int y(x) f(x) d x & \text { for continuous variables }
\end{array}
$$

Any moment $\mathrm{m}_{k}\left(\mathrm{X}_{0}\right)$ can be treated as an expected value of a function $Y(X)=\left(X-x_{0}\right)^{k}$

$$
m_{k}\left(x_{0}\right) \equiv \int\left(x-x_{0}\right)^{k} f(x) d x=E\left(\left(x-x_{0}\right)^{k}\right)
$$

## Variance

VARIANCE (dispersion) symbols: $\sigma^{2}(\mathrm{X}), \operatorname{var}(\mathrm{X}), \mathrm{V}(\mathrm{X}), \mathrm{D}(\mathrm{X})$. Standard deviation $\sigma(x)$

$$
\sigma^{2}(X) \equiv \sum_{i} p_{i}\left(x_{i}-E(X)\right)^{2} \quad \text { for discrete variables }
$$

$\sigma^{2}(X) \equiv \int f(x)\left(x-E(X)^{2} d x \quad\right.$ for continuous variables
Variance (or the standard deviation) is a measure of scatter of random variables around the expected value.

$$
\sigma^{2}(X)=E\left(X^{2}\right)-E^{2}(X)
$$

## Properties of $\boldsymbol{\sigma}^{\mathbf{2}}(\mathbf{X})$

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Variance can be calculated using expected values only:
1.

$$
\sigma^{2}(X)=E\left(X^{2}\right)-E^{2}(X)
$$

In a consequence we get:

$$
\begin{gathered}
\sigma^{2}(\mathrm{C})=0 \\
\sigma^{2}(\mathrm{CX})=\mathrm{C}^{2} \sigma^{2}(\mathrm{X}) \\
\sigma^{2}\left(\mathrm{C}_{1} \mathrm{X}+\mathrm{C}_{2}\right)=\mathrm{C}_{1}{ }^{2} \sigma^{2}(\mathrm{X})
\end{gathered}
$$

2. For independent variables $X_{1}, X_{2}, \ldots X_{n}$

$$
\sigma^{2}\left(\sum_{i} C_{i} X_{i}\right)=\sum_{i} C_{i}^{2} \sigma^{2}(X)
$$

## UNIFORM DISTRIBUTION

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$$
\begin{gathered}
f(x)=\frac{1}{b-a} \quad \mu=E X=\frac{a+b}{2} \quad \sigma^{2}=\frac{(b-a)^{2}}{12} \\
a \leq x \leq b
\end{gathered}
$$



## Czebyszew inequality

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Interpretation of variance results from Czebyszew theorem:

$$
P(|X-E(X)| \geq a . \sigma(X)) \leq \frac{1}{a^{2}}
$$

## Theorem:

Probability of the random variable $X$ to be shifted from the expected value $E(X)$ by a-times standard deviation is smaller or equal to $1 / a^{2}$

This theorem is valid for all distributions that have a variance and the expected value. Number a is any positive real value.

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## Variance as a measure of data scatter




Big scatter of data

Smaller scatter of data

## Range as a measure of scatter

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RANGE $=x_{\max }{ }^{-} x_{\text {min }}$

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## Practical ways of calculating

## variance

Variance of n-element sample:

$$
\begin{aligned}
s^{2}= & \frac{1}{n-1} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2} \quad s^{2}=\frac{1}{n-1}\left[\sum_{i=1}^{n} x_{i}^{2}-\frac{\left(\sum_{i=1}^{n} x_{i}\right)^{2}}{n}\right] \\
& \bar{x}-\text { average }
\end{aligned}
$$

Variance of N -element population :

$$
\begin{array}{r}
\sigma^{2}=\frac{1}{N} \sum_{i=1}^{N}\left(x_{i}-\mu\right)^{2} \\
\mu-\text { expected value }
\end{array}
$$

## Practical ways of calculating standard deviation

Standard deviation of sample (or: standard uncertainty):

$$
s=\sqrt{\frac{1}{n-1} \sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}}
$$

Standard deviation (population):

$$
\sigma=\sqrt{\frac{1}{N} \sum_{i=1}^{N}\left(x_{i}-\mu\right)^{2}}
$$

## Examples of probability distributions - discrete variables

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Two-point distribution (zero-one), e.g. coin toss, head = failure $x=0$, tail $=$ success $x=1, p$ - probability of success, its distribution:

| $x_{i}$ | 0 | 1 |
| :---: | ---: | :---: |
| $p_{i}$ | $1-p$ | $p$ |

Binomial (Bernoulli)

$$
p_{k}=\binom{n}{k} \cdot p^{k}(1-p)^{n-k}, k=0,1, \ldots, n
$$

where $0<p<1 ; X=\{0,1,2, \ldots k\} k-$ number of successes when n-times sampled with replacement

For $k=1$ two-point distribution

## Binomial distribution assumptions

## Random experiment consists of $n$ Bernoulli trials:

1. Each trial is independent of others.
2. Each trial can have only two results: ,,success" and ,,failure" (binary!).
3. Probability of success $\boldsymbol{p}$ is constant.

Probability $p_{k}$ of an event that random variable $X$ will be equal to the number of $k$-successes at $n$ trials.

$$
p_{k}=\binom{n}{k} \cdot p^{k}(1-p)^{n-k}, k=0,1, \ldots, n
$$

## Pascal's triangle

$$
\text { Symbol } \quad\binom{n}{k}=\frac{n!}{(n-k)!k!}
$$

$$
\begin{aligned}
& n=0 \quad\binom{0}{0}=1 \quad(a+b)^{n}=\sum_{k=0}^{n}\binom{n}{k} a^{k} b^{n-k} \\
& n=1 \quad\binom{1}{0}=1 \quad\binom{1}{1}=1 \\
& n=2 \\
& \binom{2}{0}=1 \\
& \binom{2}{1}=2 \quad\binom{2}{2}=1
\end{aligned}
$$

## Pascal's triangle



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## Bernoulli distribution

## Example 4.6

Probability that in a company the daily use of water will not exceed a certain level is $p=3 / 4$. We monitor a use of water for 6 days.
Calculate a probability the daily use of water will not exceed the set-up limit in $0,1,2, \ldots, 6$ consecutive days, respectively.

## Data:

$$
p=\frac{3}{4} \quad q=\frac{1}{4} \quad N=6 \quad k=0,1, \ldots, 6
$$

## Bernoulli distribution

$$
\begin{array}{ll}
k=0 & P(k=0)=\binom{6}{0} \cdot\left(\frac{3}{4}\right)^{0} \cdot\left(\frac{1}{4}\right)^{6} \\
k=1 & P(k=1)=\binom{6}{1} \cdot\left(\frac{3}{4}\right)^{1} \cdot\left(\frac{1}{4}\right)^{5} \\
k=2 & P(k=2)=\binom{6}{2} \cdot\left(\frac{3}{4}\right)^{2} \cdot\left(\frac{1}{4}\right)^{4} \\
k=3 & P(k=3)=\binom{6}{3} \cdot\left(\frac{3}{4}\right)^{3} \cdot\left(\frac{1}{4}\right)^{3} \\
k=4 & P(k=5)=\binom{6}{5} \cdot\left(\frac{3}{4}\right)^{5} \cdot\left(\frac{1}{4}\right)^{1} \\
k=5 & P(k=6)=\binom{6}{6} \cdot\left(\frac{3}{4}\right)^{6} \cdot\left(\frac{1}{4}\right)^{0}
\end{array}
$$

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$$
\begin{array}{ll}
k=0 & P(0)=1 \cdot 1 \cdot \frac{1}{4^{6}} \cong 0.00024 \\
k=1 & P(1)=6 \cdot \frac{3}{4} \cdot \frac{1}{4^{5}}=\frac{6 \cdot 3}{4^{6}}=18 \cdot P(0) \cong 0.004 \\
k=2 & P(2)=15 \cdot\left(\frac{3}{4}\right)^{2} \cdot \frac{1}{4^{4}}=\frac{15 \cdot 9}{4^{6}}=135 \cdot P(0) \cong 0.033 \\
k=3 & P(3)=20 \cdot\left(\frac{3}{4}\right)^{3} \cdot \frac{1}{4^{3}}=\frac{20 \cdot 9 \cdot 3}{4^{6}}=540 \cdot P(0) \cong 0.132 \\
k=4 & P(4)=15 \cdot\left(\frac{3}{4}\right)^{4} \cdot \frac{1}{4^{2}}=\frac{15 \cdot 9 \cdot 9}{4^{6}}=1215 \cdot P(0) \cong 0.297 \\
k=5 & P(5)=6 \cdot\left(\frac{3}{4}\right)^{5} \cdot \frac{1}{4^{1}}=\frac{6 \cdot 9 \cdot 9 \cdot 3}{4^{6}}=1458 \cdot P(0) \cong 0.356 \\
k=6 & P(6)=1 \cdot\left(\frac{3}{4}\right)^{6} \cdot \frac{1}{4^{0}}=\frac{9 \cdot 9 \cdot 9}{4^{6}}=729 \cdot P(0) \cong 0.178
\end{array}
$$

## Bernoulli distribution

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Maximum for $\mathrm{k}=5$


## Bernoulli distribution

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Expected value

$$
E(X)=\mu=n p
$$

Variance

$$
V(X)=\sigma^{2}=n p(1-p)
$$

## Errors in transmission

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## Example 4.7

Digital channel of information transfer is prone to errors in single bits. Assume that the probability of single bit error is $\mathrm{p}=0.1$

Consecutive errors in transmissions are independent. Let $X$ denote the random variable, of values equal to the number of bits in error, in a sequence of 4 bits.

E - bit error, O - no error
OEOE corresponds to $X=2$; for EEOO $-X=2$ (order does not matter)

## Errors in transmission

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## Example 4.7 cd

For $\mathrm{X}=2$ we get the following results:
\{EEOO, EOEO, EOOE, OEEO, OEOE, OOEE\}

What is a probability of $P(X=2)$, i.e., two bits will be sent with error?

Events are independent, thus
$P(E E O O)=P(E) P(E) P(O) P(O)=(0.1)^{2}(0.9)^{2}=0.0081$
Events are mutually exhaustive and have the same probability, hence
$P(X=2)=6 P(E E O O)=6(0.1)^{2}(0.9)^{2}=6(0.0081)=0.0486$

## Errors in transmission

Example 4.7 continued

$$
\binom{4}{2}=\frac{4!}{(2)!2!}=6
$$

Therefore, $\mathrm{P}(\mathrm{X}=2)=6(0.1)^{2}(0.9)^{2}$ is given by Bernoulli distribution

$$
P(X=x)=\binom{4}{x} \cdot p^{x}(1-p)^{4-x}, x=0,1,2,3,4, p=0.1
$$

$$
\begin{aligned}
& P(X=0)=0,6561 \\
& P(X=1)=0,2916 \\
& P(X=2)=0,0486 \\
& P(X=3)=0,0036 \\
& P(X=4)=0,0001
\end{aligned}
$$



## Poisson's distribution

We introduce a parameter $\lambda=p n(E(X)=\lambda)$

$$
P(X=x)=\binom{n}{x} \cdot p^{x}(1-p)^{n-x}=\binom{n}{x}\left(\frac{\lambda}{n}\right)^{x}\left(1-\frac{\lambda}{n}\right)^{n-x}
$$

Let us assume that n increases while p decreases, but $\lambda=\mathrm{pn}$ remains constant. Bernoulli distribution changes to Poisson's distribution.

$$
\lim _{n \rightarrow \infty} P(X=x)=\lim _{n \rightarrow \infty}\binom{n}{x}\left(\frac{\lambda}{n}\right)^{x}\left(1-\frac{\lambda}{n}\right)^{n-x}=\frac{e^{-\lambda} \lambda^{x}}{x!}
$$

## Poisson's distribution

It is one of the rare cases where expected value equals to variance:

$$
E(X)=n p=\lambda
$$

Why?

$$
V(X)=\sigma^{2}=\lim _{n \rightarrow \infty, p \rightarrow 0}\left(n p-n p^{2}\right)=n p=\lambda
$$

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## Poisson's distribution


( $x$ - integer, infinite; $x \geq 0$ ) For big $n$ Bernoulli distribution resembles Poisson's distribution

## Normal distribution (Gaussian)

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The most widely used model for the distribution of random variable is a normal distribution.
Central limit theorem formulated in 1733 by De Moivre
Whenever a random experiment is replicated, the random variable that equals the average (or total) result over the replicas tends to have a normal distribution as the number of replicas becomes large.

## Normal distribution (Gaussian)

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A random variable $X$ with probability density function $f(x)$ :

$$
f(x)=\frac{1}{\sigma \sqrt{2 \pi}} \exp \left[-\frac{\left(x-\mu^{2}\right.}{2 \sigma^{2}}\right], \text { where }-\infty<x<+\infty
$$

is a normal random variable with two parameters:

$$
-\infty<\mu<+\infty, \quad \sigma>1
$$

We can show that $E(X)=\mu$ and $V(X)=\sigma^{2}$

Notation $N(\mu, \sigma)$ is used to denote this distribution

## Normal distribution (Gaussian)

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Expected value, maximum of density probability (mode) and median overlap ( $x=\mu$ ). Symmetric curve (Gaussian curve is bell shaped).

Variance is a measure of the width of distribution. At $x=+\sigma$ and $x=-\sigma$ there are the inflection points of $\mathrm{N}(0, \sigma)$.

## Normal distribution (Gaussian)

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Is used in experimental physics and describes distribution of random errors. Standard deviation $\sigma$ is a measure of random uncertainty. Measurements with larger $\sigma$ correspond to bigger scatter of data around the average value and thus have less precision.

## Standard normal distribution

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A normal random variable $Z$ with probability density $N(z)$ :

$$
N(z)=\frac{1}{\sqrt{2 \pi}} \exp \left[-\frac{z^{2}}{2}\right] \text {, where }-\infty<z<+\infty
$$

is called a standard normal random variable

$$
\mathrm{N}(0,1) \quad E(Z)=0, \quad V(Z)=1
$$

Definition of standard normal variable

$$
Z=\frac{X-\mu}{\sigma}
$$

## Standard normal distribution

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Advantages of standardization:

- Tables of values of probability density and CDF can be constructed for $N(0,1)$. A new variable of the $N(\mu, \sigma)$ distribution can be created by a simple transformation $\mathrm{X}=\sigma^{*} \mathrm{Z}+\mu$
- By standardization we shift all original random variables to the region close to zero and we rescale the x-axis. The unit changes to standard deviation. Therefore, we can compare different distribution.


## Calculations of probability (Gaussian distribution)

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$$
\begin{aligned}
& \mathrm{P}(\mu-\sigma<X<\mu+\sigma)=0,6827 \text { (about } 2 / 3 \text { of results) } \\
& \mathrm{P}(\mu-2 \sigma<X<\mu+2 \sigma)=0,9545 \\
& \mathrm{P}(\mu-2 \sigma<X<\mu+2 \sigma)=0,9973 \text { (almost all) }
\end{aligned}
$$

